

A MULTIDIMENSIONAL ANALOGUE OF THE ARCSINE LAW FOR THE NUMBER OF POSITIVE TERMS IN A RANDOM WALK

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ABSTRACT. Consider a random walk $S_i = \xi_1 + \dots + \xi_i$, $i \in \mathbb{N}$, whose increments ξ_1, ξ_2, \dots are independent identically distributed random vectors in \mathbb{R}^d such that ξ_1 has the same law as $-\xi_1$ and $\mathbb{P}[\xi_1 \in H] = 0$ for every affine hyperplane $H \subset \mathbb{R}^d$. Our main result is the distribution-free formula

$$\mathbb{E} \left[\sum_{1 \leq i_1 < \dots < i_k \leq n} \mathbb{1}_{\{0 \notin \text{Conv}(S_{i_1}, \dots, S_{i_k})\}} \right] = 2 \binom{n}{k} \frac{B(k, d-1) + B(k, d-3) + \dots}{2^k k!},$$

where the $B(k, j)$'s are defined by their generating function $(t+1)(t+3)\dots(t+2k-1) = \sum_{j=0}^k B(k, j)t^j$. The expected number of k -tuples above admits the following geometric interpretation: it is the expected number of k -dimensional faces of a randomly and uniformly sampled open Weyl chamber of type B_n that are not intersected by a generic linear subspace $L \subset \mathbb{R}^n$ of codimension d . The case $d = 1$ turns out to be equivalent to the classical discrete arcsine law for the number of positive terms in a one-dimensional random walk with continuous symmetric distribution of increments. We also prove similar results for random bridges with no central symmetry assumption required.

1. INTRODUCTION AND MAIN RESULTS

1.1. Introduction. Let $\xi_1, \dots, \xi_n \in \mathbb{R}^1$ be i.i.d. random variables with continuous, symmetric distribution, that is, $\mathbb{P}[\xi_1 = x] = 0$ and $\mathbb{P}[\xi_1 < -x] = \mathbb{P}[\xi_1 > x]$ for all $x \in \mathbb{R}$. Consider the one-dimensional *random walk* $S_i := \xi_1 + \dots + \xi_i$, $1 \leq i \leq n$. We are interested in the random variable

$$N_n = \sum_{i=1}^n \mathbb{1}_{\{S_i > 0\}} \tag{1}$$

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counting the number of positive terms in the random walk. The classical *discrete arcsine law* due to Sparre Andersen [8, Theorem 1 and Eq. (8)] and [9, Theorem 4] states that

$$\mathbb{P}[N_n = m] = \frac{1}{2^{2n}} \binom{2m}{m} \binom{2n-2m}{n-m}, \quad m = 0, \dots, n. \quad (2)$$

Passing to the limit gives

$$\lim_{n \rightarrow \infty} \mathbb{P} \left[\frac{N_n}{n} \leq x \right] = \frac{2}{\pi} \arcsin \sqrt{x}, \quad 0 \leq x \leq 1, \quad (3)$$

which justifies the name of the law in (2).

A discussion of this remarkable result, together with a simplified, purely combinatorial proof, can be found in Feller's book [5, Vol II, Section XII.8]. Observe that formula (2) is distribution-free since the values on the right-hand side do not depend on the distribution of ξ_1 as long as this distribution is continuous and symmetric.

The aim of the present paper is to obtain a multidimensional generalization of this result to random walks in \mathbb{R}^d . Let us start with a special case. For $m = 0$ and $m = n$, formula (2) provides an expression for the so-called *persistence probability*

$$\mathbb{P}[N_n = 0] = \mathbb{P}[S_1 < 0, \dots, S_n < 0] = \mathbb{P}[S_1 > 0, \dots, S_n > 0] = \mathbb{P}[N_n = n] = \frac{1}{2^{2n}} \binom{2n}{n}. \quad (4)$$

This formula has been generalized to the d -dimensional case in the following way [6]. Consider a d -dimensional random walk $S_i = \xi_1 + \dots + \xi_i$, $1 \leq i \leq n$, whose increments ξ_1, ξ_2, \dots are i.i.d. random vectors in \mathbb{R}^d , $d \in \mathbb{N}$, with centrally symmetric distribution, that is $\xi_1 \stackrel{d}{=} -\xi_1$. Denote by $\text{Conv}(x_1, \dots, x_k)$ the *convex hull* of any k points $x_1, \dots, x_k \in \mathbb{R}^d$, that is

$$\text{Conv}(x_1, \dots, x_k) = \{\alpha_1 x_1 + \dots + \alpha_k x_k : \alpha_1, \dots, \alpha_k \geq 0, \alpha_1 + \dots + \alpha_k = 1\},$$

and consider the *non-absorption probability* $\mathbb{P}[0 \notin \text{Conv}(S_1, \dots, S_n)]$. If $d = 1$, this probability equals $2\mathbb{P}[N_n = 0]$ because $0 \notin \text{Conv}(S_1, \dots, S_n)$ if and only if either $S_1 > 0, \dots, S_n > 0$ or $S_1 < 0, \dots, S_n < 0$, and the probabilities of these two events are equal by the symmetry assumption. For general $d \in \mathbb{N}$, a distribution-free formula for the non-absorption probability has been obtained in [6] and will be recalled in Example 1.4 below. In the special case $d = 1$, this formula reduces to (4).

When searching for a multidimensional generalization of (2) for general $0 \leq m \leq n$, the basic question is how to define positivity in \mathbb{R}^d . One possible approach is to declare a vector positive if all of its components are positive. This leads to the question on how much time a random walk or a Brownian motion spends in the positive orthant \mathbb{R}_+^d ; see the work of Bingham and Doney [3] which contains a review on the higher-dimensional analogues of the arcsine law of this type. In the present paper we choose a different, coordinate-free approach.

The main idea is that instead of looking at the random variable N_n itself, we shall find an appropriate generalization of its factorial moments. The advantage of working with factorial moments instead of the usual power moments will become evident later.

Observe that we can rewrite (2) as follows:

$$\mathbb{E} \left[\binom{N_n}{k} \right] = \sum_{m=k}^n \frac{1}{2^{2m}} \binom{2m}{m} \binom{2n-2m}{n-m} \binom{m}{k} = \frac{1}{2^{2k}} \binom{2k}{k} \binom{n}{k}, \quad k = 0, 1, \dots, n. \quad (5)$$

We omit a direct proof of the second equality because it will be recovered as a special case of Theorem 1.2 presented below. Note that since N_n takes values in $\{0, \dots, n\}$, the factorial moments in (5) determine the law of N_n uniquely and therefore statements (2) and (5) are indeed equivalent. In fact, (5) can be viewed as a system of $n + 1$ linear equations in the unknowns $\mathbb{P}[N_n = i]$, $0 \leq i \leq n$, with a non-degenerate upper triangular matrix.

We shall give our multidimensional arcsine law in the form of a d -dimensional version of (5). Our main result is in showing that this statement admits an interpretation in terms of an equivalent geometric problem concerning Weyl chambers intersected by a generic linear subspace; see Theorem 2.1. This geometric interpretation seems to be new even in the one-dimensional case of the discrete arcsine law while the proofs of the arcsine law given in [8], [9], and [5] are purely combinatorial.

Moreover, in the special one-dimensional case there is a different (and new) geometric interpretation of the discrete arcsine law (2) itself. Let us describe it. Consider the simplex

$$\{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : 1 \geq \beta_1 \geq \beta_2 \geq \dots \geq \beta_n \geq 0\}$$

or, equivalently, the convex hull of the following $n + 1$ points:

$$(0, 0, \dots, 0), (1, 0, \dots, 0), (1, 1, \dots, 0), \dots, (1, 1, \dots, 1).$$

There are $2^n n!$ isometric simplices obtained by applying to the above simplex orthogonal transformations of \mathbb{R}^n that permute the coordinates and change their signs. In other words, these simplices are the closed Weyl chambers of type B_n intersected with the cube $[-1, 1]^n$. Their union is exactly $[-1, 1]^n$, and the interiors of the simplices are disjoint.

Let $H \subset \mathbb{R}^n$ be any generic open half-space with the boundary passing through the origin. It will be shown that the discrete arcsine probability in (2) equals the fraction of the above simplices with exactly m vertices lying in H ; see Corollary 2.2 in Section 2 below. The asymptotic arcsine law (3) interprets as follows: for any fixed $x \in [0, 1]$, the relative fraction of the simplices having at most xn vertices in H tends to $\frac{2}{\pi} \arcsin \sqrt{x}$ as $n \rightarrow \infty$.

1.2. Arcsine law for random walks. We shall give a generalization of (5) to random walks in \mathbb{R}^d . Let ξ_1, \dots, ξ_n be random d -dimensional vectors. To avoid trivialities, we always assume that $n \geq d + 1$. The d -dimensional random walk $(S_i)_{i=1}^n$ with increments ξ_1, \dots, ξ_n is defined by

$$S_i := \xi_1 + \dots + \xi_i, \quad 1 \leq i \leq n.$$

We impose the following assumptions on the increments ξ_1, \dots, ξ_n :

(±Ex) *Symmetric exchangeability:* For every permutation σ of the set $\{1, \dots, n\}$ and every $\varepsilon_1, \dots, \varepsilon_n \in \{-1, +1\}$ there is the distributional equality

$$(\xi_1, \dots, \xi_n) \stackrel{d}{=} (\varepsilon_1 \xi_{\sigma(1)}, \dots, \varepsilon_n \xi_{\sigma(n)}).$$

(GP) *General position:* For every $1 \leq i_1 < \dots < i_d \leq n$, the probability that the vectors S_{i_1}, \dots, S_{i_d} are linearly dependent, is 0.

Remark 1.1. If ξ_1, ξ_2, \dots are independent identically distributed in \mathbb{R}^d and such that ξ_1 has the same distribution as $-\xi_1$, then $(\pm Ex)$ is satisfied and the following conditions are equivalent:

- (i) (GP) holds for all $n \geq d + 1$;
- (ii) for every affine hyperplane $H \subset \mathbb{R}^d$ we have $\mathbb{P}[\xi_1 \in H] = 0$;
- (iii) for every hyperplane $H_0 \subset \mathbb{R}^d$ passing through the origin and every $i \in \mathbb{N}$, we have $\mathbb{P}[S_i \in H_0] = 0$.

This statement is proved in [6, Proposition 2.5] (which does assume that $\xi_1 \stackrel{d}{=} -\xi_1$ but does not state this explicitly in the published version).

For $1 \leq k \leq n$ we are interested in the random variable equal to half the number of polytopes of the form $\text{Conv}(S_{i_1}, \dots, S_{i_k})$ that do not contain the origin:

$$M_{n,k}^{(d)} = \frac{1}{2} \sum_{1 \leq i_1 < \dots < i_k \leq n} \mathbb{1}_{\{0 \notin \text{Conv}(S_{i_1}, \dots, S_{i_k})\}}. \quad (6)$$

In the one-dimensional case $d = 1$ the convex hull of S_{i_1}, \dots, S_{i_k} does not contain the origin if and only if the numbers S_{i_1}, \dots, S_{i_k} have the same sign, whence

$$M_{n,k}^{(1)} = \frac{1}{2} \binom{N_n}{k} + \frac{1}{2} \binom{n - N_n}{k} \quad \text{a.s.}$$

Here we used that $\mathbb{P}[S_i = 0] = 0$, $1 \leq i \leq n$, which holds by the general position assumption (GP). Using the fact that N_n has the same distribution as $n - N_n$, which is a consequence of assumption $(\pm Ex)$, we deduce that

$$\mathbb{E}M_{n,k}^{(1)} = \mathbb{E} \left[\binom{N_n}{k} \right]. \quad (7)$$

Therefore, we can view $\mathbb{E}M_{n,k}^{(d)}$ as a d -dimensional generalization of $\mathbb{E} \left[\binom{N_n}{k} \right]$. Our main result generalizes (5) to arbitrary dimension as follows.

Theorem 1.2. *Consider a random walk $(S_i)_{i=1}^n$ in \mathbb{R}^d , $n \geq d + 1$, with increments ξ_1, \dots, ξ_n satisfying assumptions $(\pm Ex)$ and (GP). For every $k = 1, \dots, n$ we have*

$$\mathbb{E}M_{n,k}^{(d)} = \binom{n}{k} \frac{B(k, d-1) + B(k, d-3) + \dots}{2^k k!} = \binom{n}{k} \mathbb{E}M_{k,k}^{(d)}, \quad (8)$$

where the $B(k, j)$'s are defined by their generating function

$$(t+1)(t+3) \dots (t+2k-1) = \sum_{j=0}^k B(k, j) t^j. \quad (9)$$

We put $B(k, j) = 0$ for $j < 0$ and $j > k$ so that the sum in (8) has only finitely many non-zero terms.

Example 1.3. For $d = 1$ Theorem 1.2 reduces to (5) in view of $B(k, 0) = (2k-1)!!$:

$$\mathbb{E} \left[\binom{N_n}{k} \right] = \mathbb{E}M_{n,k}^{(1)} = \binom{n}{k} \frac{(2k-1)!!}{2^k k!} = \frac{1}{2^{2k}} \binom{2k}{k} \binom{n}{k}.$$

Note in passing that this proves the second equality in (5).

Example 1.4. In the case $k = n$ Theorem 1.2 provides a formula for the non-absorption probability

$$\mathbb{P}[0 \notin \text{Conv}(S_1, \dots, S_n)] = \frac{2(B(n, d-1) + B(n, d-3) + \dots)}{2^n n!}. \quad (10)$$

This formula was obtained in [6].

Remark 1.5. For the number of polytopes of the form $\text{Conv}(S_{i_1}, \dots, S_{i_k})$ containing the origin we have the formula

$$\mathbb{E} \left[\sum_{1 \leq i_1 < \dots < i_k \leq n} \mathbb{1}_{\{0 \in \text{Conv}(S_{i_1}, \dots, S_{i_k})\}} \right] = 2 \binom{n}{k} \frac{B(k, d+1) + B(k, d+3) + \dots}{2^k k!} \quad (11)$$

which follows from (8) and the identities

$$B(k, 2) + B(k, 4) + \dots = B(k, 1) + B(k, 3) + \dots = 2^{k-1} k!.$$

To prove these identities, take $t = \pm 1$ in (9). Note that for $1 \leq k \leq d$ both sides of (11) vanish (the left hand-side vanishes by assumption (GP)).

Let us now pass to the large n limit. The classical arcsine law [4, 8] for the number of positive terms in a one-dimensional random walk (whose increments are symmetrically distributed or have zero mean and finite positive variance, the case not considered in our paper) can be stated as follows:

$$\lim_{n \rightarrow \infty} \mathbb{P} \left[\frac{N_n}{n} \leq x \right] = \frac{2}{\pi} \arcsin \sqrt{x}, \quad 0 \leq x \leq 1. \quad (12)$$

In terms of moments (which fully define any distribution concentrated on $[0, 1]$), this can equivalently be written as

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{N_n^k}{n^k} \right] = \frac{1}{2^{2k}} \binom{2k}{k}, \quad k \in \mathbb{N}. \quad (13)$$

Note that

$$\mathbb{E} N_n^k = k! \mathbb{E} \left[\binom{N_n}{k} \right] (1 + o(1)), \quad n \rightarrow \infty, \quad (14)$$

which together with (7) implies that (13) is equivalent to

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{k! M_{n,k}^{(1)}}{n^k} \right] = \frac{1}{2^{2k}} \binom{2k}{k}, \quad k \in \mathbb{N}.$$

From Theorem 1.2 we obtain the following d -dimensional generalization of (14):

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{k! M_{n,k}^{(d)}}{n^k} \right] = \frac{B(k, d-1) + B(k, d-3) + \dots}{2^k k!} = \mathbb{E} M_{k,k}^{(d)}, \quad k \in \mathbb{N}. \quad (15)$$

Remark 1.6. Thus, in the d -dimensional setting, the sequence in (15) is analogous to the sequence of moments of the arcsine distribution. However, unlike the one-dimensional case, for $d \geq 2$ this sequence does not correspond to a distribution on $[0, 1]$ because both the first and the second moments of such hypothetical distribution should be $\frac{1}{2}$. On the other hand, it allows the following natural interpretation.

Let $W^{(d)}(t)$, $t \in [0, 1]$, be a standard Brownian motion taking values in \mathbb{R}^d . Consider the random variable

$$M_{\infty, k}^{(d)} := \frac{1}{2} \int_{0 < t_1 < \dots < t_k < 1} \mathbb{1}_{\{0 \notin \text{Conv}(W^{(d)}(t_1), \dots, W^{(d)}(t_k))\}} dt_1 \dots dt_k.$$

Note that if U_1, \dots, U_k are i.i.d. random variables distributed uniformly on $[0, 1]$ and independent of $W^{(d)}$, then

$$\mathbb{E}M_{\infty, k}^{(d)} = \frac{1}{2 \cdot k!} \mathbb{P}[0 \notin \text{Conv}(W^{(d)}(U_1), \dots, W^{(d)}(U_k))].$$

Let k and d be fixed, while $n \rightarrow \infty$. Using Donsker's invariance principle, it is possible to show that $\frac{1}{n^k} M_{n, k}^{(d)}$ converges weakly (together with all moments) to $M_{\infty, k}^{(d)}$. Hence, by (15),

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{k! M_{n, k}^{(d)}}{n^k} \right] = k! \mathbb{E}M_{\infty, k}^{(d)} = \mathbb{E}M_{k, k}^{(d)} = \frac{1}{2} \mathbb{P}[0 \notin \text{Conv}(W^{(d)}(U_1), \dots, W^{(d)}(U_k))]. \quad (16)$$

The same result may be obtained by observing that the increments of the sequence $W^{(d)}(U_{(1)}), \dots, W^{(d)}(U_{(k)})$, where $U_{(1)}, \dots, U_{(k)}$ are the order statistics of U_1, \dots, U_k , are exchangeable, and applying Example 1.4 directly. This explains why $\mathbb{E}M_{k, k}^{(d)}$ appears in (16). We also note that the equality $\mathbb{E}M_{k, k}^{(1)} = k! \mathbb{E}M_{\infty, k}^{(1)}$ implies directly that $\mathbb{E}M_{k, k}^{(1)}$ is the k -th moment of the arcsine law. This is easily seen from the definition of $M_{\infty, k}^{(1)}$ and the following arcsine law for the Brownian motion:

$$\mathbb{P} \left[\int_0^1 \mathbb{1}_{\{W^{(1)}(t) > 0\}} dt \leq x \right] = \frac{2}{\pi} \arcsin \sqrt{x}, \quad 0 \leq x \leq 1.$$

Remark 1.7. Theorem 1.2 shows that the *expectation* of $M_{n, k}^{(d)}$ does not depend on the distribution of increments of the random walk. One may ask whether the *distribution* of $M_{n, k}^{(d)}$ has the same property. Our simulations gave a strong evidence against this conjecture.

We shall give two proofs of Theorem 1.2. The first proof deduces Theorem 1.2 from a geometric result on the number of k -faces of a random Weyl chamber that are intersected by a linear subspace. This result, which is of an independent interest, will be stated in Theorem 2.1, Section 2. The second proof, given in Section 3.1, is based on (10). Both proofs strongly rely on the ideas and results of [6]. The second proof is shorter but it does not allow any interpretation of Theorem 1.2. In the next section, a similar statement for random bridges is formulated in Theorem 1.8 (which will be proved in Section 2.6).

1.3. Uniform law for random bridges. Similar results can be obtained for random bridges which are essentially random walks required to return to the origin after n steps. Let ξ_1, \dots, ξ_n be random vectors in \mathbb{R}^d , where the reader may always assume that $n \geq d + 2$ to avoid trivialities. We define the partial sums $(S_i)_{i=1}^n$ by

$$S_i := \xi_1 + \dots + \xi_i, \quad 1 \leq i \leq n,$$

and impose the following assumptions on the increments ξ_1, \dots, ξ_n :

(Br) *Bridge property:* $S_n = \xi_1 + \dots + \xi_n = 0$ a.s.

(Ex) *Exchangeability:* For every permutation σ of the set $\{1, \dots, n\}$ we have the distributional equality

$$(\xi_{\sigma(1)}, \dots, \xi_{\sigma(n)}) \stackrel{d}{=} (\xi_1, \dots, \xi_n).$$

(GP') *General position:* For every $1 \leq i_1 < \dots < i_d \leq n - 1$, the probability that the vectors S_{i_1}, \dots, S_{i_d} are linearly dependent, is 0.

The stochastic process $(S_i)_{i=1}^n$ is called a *random bridge*. Note that assumption (Ex) does not require invariance with respect to sign changes and thus is weaker than $(\pm\text{Ex})$.

For $d = 1$ the distribution of the random variable N_n counting the number of positive terms among S_1, \dots, S_{n-1} is discrete uniform on $\{0, \dots, n - 1\}$ by a result of Sparre Andersen [10, Corollary 2]. That is,

$$\mathbb{P}[N_n = m] = \frac{1}{n}, \quad m = 0, \dots, n - 1. \quad (17)$$

Alternatively, this formula follows easily from [11, Theorem 2.1]. In terms of factorial moments, (17) can be stated as follows:

$$\mathbb{E} \left[\binom{N_n}{k} \right] = \frac{1}{n} \sum_{m=k}^{n-1} \binom{m}{k} = \frac{1}{k+1} \binom{n-1}{k}, \quad k = 0, \dots, n - 1. \quad (18)$$

The second equality in (18) follows easily by induction over n . Note that (17) and (18) are equivalent similarly to the case of random walks we seen above.

To state a d -dimensional generalization of (18), we consider a slight modification of $M_{n,k}^{(d)}$, namely

$$M_{n,k}^{(d)} = \frac{1}{2} \sum_{1 \leq i_1 < \dots < i_k \leq n-1} \mathbb{1}_{\{0 \notin \text{Conv}(S_{i_1}, \dots, S_{i_k})\}}. \quad (19)$$

We excluded the case $i_k = n$ because the corresponding convex hulls would contain 0 by the assumption $S_n = 0$ a.s. The following is our main result for random bridges.

Theorem 1.8. *Consider a random bridge $(S_i)_{i=1}^n$ in \mathbb{R}^d , $n \geq d + 2$, whose increments ξ_1, \dots, ξ_n satisfy assumptions (Br), (Ex), (GP'). For all $k = 1, \dots, n - 1$ we have*

$$\mathbb{E} M_{n,k}^{(d)} = \frac{1}{(k+1)!} \binom{n-1}{k} \left(\binom{k+1}{d} + \binom{k+1}{d-2} + \dots \right) = \binom{n-1}{k} \mathbb{E} M_{k+1,k}^{(d)}, \quad (20)$$

where $\left[\begin{smallmatrix} k+1 \\ 1 \end{smallmatrix} \right], \dots, \left[\begin{smallmatrix} k+1 \\ k+1 \end{smallmatrix} \right]$ are the Stirling numbers of the first kind defined by the formula

$$t(t+1)\dots(t+k) = \sum_{j=1}^{k+1} \left[\begin{smallmatrix} k+1 \\ j \end{smallmatrix} \right] t^j. \quad (21)$$

We use the convention $\left[\begin{smallmatrix} k+1 \\ j \end{smallmatrix} \right] = 0$ for $j < 0$ and $j > k+1$, so the sum in (20) contains a finite number of non-zero terms.

Example 1.9. In the one-dimensional case $d = 1$ we have

$$M_{n,k}^{(1)} = \frac{1}{2} \binom{N_n}{k} + \frac{1}{2} \binom{n - N_n - 1}{k} \quad \text{a.s.}$$

Theorem 1.8 yields

$$\frac{1}{2} \mathbb{E} \left[\binom{N_n}{k} + \binom{n - N_n - 1}{k} \right] = \mathbb{E} M_{n,k}^{(1)} = \frac{1}{(k+1)!} \binom{n-1}{k} \left[\begin{smallmatrix} k+1 \\ 1 \end{smallmatrix} \right] = \frac{1}{k+1} \binom{n-1}{k}$$

in view of $\left[\begin{smallmatrix} k+1 \\ 1 \end{smallmatrix} \right] = k!$. Then we recover (18) using the distributional equality $N_n \stackrel{d}{=} n-1 - N_n$, which itself follows by $(S_i)_{i=1}^n \stackrel{d}{=} (S_n - S_{n-i})_{i=1}^n \stackrel{\text{a.s.}}{=} -(S_{n-i})_{i=1}^n$, a consequence of (Ex) and (Br).

Example 1.10. In the case $k = n - 1$ Theorem 1.8 reduces to the formula for the non-absorption probability

$$\mathbb{P}[0 \notin \text{Conv}(S_1, \dots, S_{n-1})] = \frac{2}{n!} \left(\left[\begin{smallmatrix} n \\ d \end{smallmatrix} \right] + \left[\begin{smallmatrix} n \\ d-2 \end{smallmatrix} \right] + \dots \right)$$

which was obtained in [6].

As $n \rightarrow \infty$, the random variable N_n/n converges weakly to the uniform distribution on the interval $[0, 1]$, namely

$$\lim_{n \rightarrow \infty} \mathbb{P} \left[\frac{N_n}{n} \leq x \right] = x, \quad x \in [0, 1].$$

In terms of moments, we can write this as

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{N_n^k}{n^k} \right] = \frac{1}{k+1}, \quad k \in \mathbb{N},$$

which is a bridge analogue of (13) and, similarly, equivalent to

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{k! M_{n,k}^{(1)}}{n^k} \right] = \frac{1}{k+1}, \quad k \in \mathbb{N}.$$

Theorem 1.8 yields the following d -dimensional version of this relation:

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{k! M_{n,k}^{(d)}}{n^k} \right] = \frac{1}{(k+1)!} \left(\left[\begin{smallmatrix} k+1 \\ d \end{smallmatrix} \right] + \left[\begin{smallmatrix} k+1 \\ d-2 \end{smallmatrix} \right] + \dots \right) = \mathbb{E} M_{k+1,k}^{(d)}, \quad k \in \mathbb{N}.$$

Finally, there is the following analogue of (16): if $W_0^{(d)}(t)$, $t \in [0, 1]$, is a Brownian bridge in \mathbb{R}^d and U_1, \dots, U_k are i.i.d. random variables distributed uniformly on $[0, 1]$ and independent of $W_0^{(d)}$, then

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\frac{k! M_{n,k}^{(d)}}{n^k} \right] = \mathbb{E} M_{k+1,k}^{(d)} = \frac{1}{2} \mathbb{P}[0 \notin \text{Conv}(W_0^{(d)}(U_1), \dots, W_0^{(d)}(U_k))].$$

2. RELATION TO LINEAR SUBSPACES INTERSECTING WEYL CHAMBERS

In this section we prove Theorems 1.2 and 1.8 by deducing them from certain geometric results on linear subspaces intersecting Weyl chambers of types B_n and A_{n-1} , respectively. We start by recalling the necessary definitions.

2.1. The reflection group and Weyl chambers of type B_n . The *reflection group* $\mathcal{G}(B_n)$ of type B_n acts on \mathbb{R}^n by permuting the coordinates in an arbitrary way and by multiplying any number of coordinates by -1 . That is, the elements of $\mathcal{G}(B_n)$ are isometries of the form

$$g_{\varepsilon, \sigma} : \mathbb{R}^n \rightarrow \mathbb{R}^n, \quad (\beta_1, \dots, \beta_n) \mapsto (\varepsilon_1 \beta_{\sigma(1)}, \dots, \varepsilon_n \beta_{\sigma(n)}),$$

where $\sigma \in \text{Sym}(n)$ is a permutation of the set $\{1, \dots, n\}$ and $\varepsilon = (\varepsilon_1, \dots, \varepsilon_n) \in \{-1, +1\}^n$. Here we denote by $\text{Sym}(n)$ the symmetric group on the set $\{1, \dots, n\}$. The group $\mathcal{G}(B_n)$ is the symmetry group of the n -dimensional cube $[-1, 1]^n$ and the number of elements in this group is $2^n n!$.

A set $Q \subset \mathbb{R}^n$ is called a *convex cone* if for all $x, x' \in Q$ and $\alpha, \alpha' > 0$ we have $\alpha x + \alpha' x' \in Q$. We refer to [1, 2] and [7, Section 6.5] for information on convex cones and spherical convex geometry. We shall consider only *polyhedral cones*. These are defined as finite intersections of half-spaces whose boundaries pass through the origin. The faces of the cone are obtained by replacing in the above definition some of the half-spaces by their boundaries and taking the intersection. The *fundamental Weyl chamber of type B_n* is the convex cone given by

$$\mathcal{C}(B_n) = \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : 0 < \beta_1 < \beta_2 < \dots < \beta_n\}.$$

This is a fundamental domain for $\mathcal{G}(B_n)$, meaning that the cones $g\mathcal{C}(B_n)$, $g \in \mathcal{G}(B_n)$, are disjoint and their closures (which will be called *closed Weyl chambers* or, without any risk of confusion, simply *Weyl chambers*) cover the whole \mathbb{R}^n . Thus, the closed Weyl chambers are the convex cones given by

$$C_{\varepsilon, \sigma}^B := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \varepsilon_1 \beta_{\sigma(1)} \geq \varepsilon_2 \beta_{\sigma(2)} \geq \dots \geq \varepsilon_n \beta_{\sigma(n)} \geq 0\},$$

where $\varepsilon \in \{-1, +1\}^n$, $\sigma \in \text{Sym}(n)$. The terms are required to be non-increasing rather than increasing for convenience of proofs, and the superscript B refers to the type of the chambers. We denote by $\mathcal{F}_k(Q)$ the set of all (closed) k -dimensional faces of a convex cone Q . For $1 \leq k \leq n$, the k -dimensional faces of $C_{\varepsilon, \sigma}^B$ are indexed by collections $1 \leq i_1 < \dots < i_k \leq n$

and have the form

$$\begin{aligned} C_{\varepsilon,\sigma}^B(i_1, \dots, i_k) &:= \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \varepsilon_1 \beta_{\sigma(1)} = \dots = \varepsilon_{i_1} \beta_{\sigma(i_1)} \\ &\geq \varepsilon_{i_1+1} \beta_{\sigma(i_1+1)} = \dots = \varepsilon_{i_2} \beta_{\sigma(i_2)} \geq \dots \geq \varepsilon_{i_{k-1}+1} \beta_{\sigma(i_{k-1}+1)} = \dots = \varepsilon_{i_k} \beta_{\sigma(i_k)} \\ &\geq \beta_{\sigma(i_k+1)} = \dots = \beta_{\sigma(n)} = 0\}. \end{aligned} \quad (22)$$

In the case $i_k = n$, no β_i 's are required to be 0. In particular, $\#\mathcal{F}_k(C_{\varepsilon,\sigma}^B) = \binom{n}{k}$.

A *hyperplane arrangement* is a finite set of (distinct) hyperplanes in \mathbb{R}^n . The *reflection arrangement* $\mathcal{A}(B_n)$ of type B_n consists of the hyperplanes

$$\{\beta_i = \beta_j\} \quad (1 \leq i < j \leq n), \quad \{\beta_i = -\beta_j\} \quad (1 \leq i < j \leq n), \quad \{\beta_i = 0\} \quad (1 \leq i \leq n). \quad (23)$$

The name is due to the fact that reflections with respect to these hyperplanes generate the group $\mathcal{G}(B_n)$. The *lattice* $\mathcal{L}(B_n)$ generated by the reflection arrangement of type B_n consists of linear subspaces of \mathbb{R}^n which can be represented as intersections of the hyperplanes (23). We say that a linear subspace $L \subset \mathbb{R}^n$ of codimension d is *in general position* with respect to the reflection arrangement if for every linear subspace $K \in \mathcal{L}(B_n)$ we have

$$\dim(L \cap K) = \begin{cases} \dim K - d, & \text{if } \dim K \geq d, \\ 0, & \text{if } \dim K \leq d. \end{cases} \quad (24)$$

2.2. Subspaces intersecting faces of Weyl chambers of type B_n . The next theorem, which is the main result of the present section, will be shown to imply Theorem 1.2.

Theorem 2.1. *Let $L \subset \mathbb{R}^n$ be a deterministic linear subspace of codimension d in general position with respect to the reflection arrangement (23) of type B_n . Let Q be sampled randomly and uniformly among the $2^n n!$ closed Weyl chambers $C_{\varepsilon,\sigma}^B$ of type B_n . Then the expected number of k -dimensional faces of Q intersected by L in a trivial way is given by*

$$\begin{aligned} \mathbb{E} \left[\sum_{F \in \mathcal{F}_k(Q)} \mathbb{1}_{\{F \cap L = \{0\}\}} \right] &\stackrel{\text{def}}{=} \frac{1}{2^n n!} \sum_{\varepsilon \in \{-1, +1\}^n} \sum_{\sigma \in \text{Sym}(n)} \sum_{F \in \mathcal{F}_k(C_{\varepsilon,\sigma}^B)} \mathbb{1}_{\{F \cap L = \{0\}\}} \\ &= 2 \binom{n}{k} \frac{B(k, d-1) + B(k, d-3) + \dots}{2^k k!}, \end{aligned}$$

where the $B(k, j)$'s are defined in Theorem 1.2.

Here, we say that Q and L intersect in a *trivial way* if $Q \cap L = \{0\}$.

Corollary 2.2. *Let $d = 1$ (so L is a hyperplane) and let H be either of the open half-spaces with the boundary L . Then the number of vertices V_n of the random simplex $Q \cap [-1, 1]^n$ lying in H follows the discrete arcsine distribution (2).*

Proof of Corollary 2.2. The number of the k -dimensional faces of Q intersected by L in a trivial way equals $\binom{V_n}{k} + \binom{n-V_n}{k}$. By taking the expectations and using the distributional identity $V_n \stackrel{d}{=} n - V_n$ and the result of Theorem 2.1, we conclude that the random variables V_n and N_n have the same factorial moments given in (4). As we argued in the Introduction, these are factorial moments of the discrete arcsine distribution, which is the unique distribution on

$\{0, \dots, n\}$ with the given factorial moments. Therefore V_n and N_n have the same distribution. \square

2.3. Proof of Theorem 1.2 given Theorem 2.1. We shall need a short notation for one of the closed Weyl chambers:

$$C^B := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \beta_1 \geq \beta_2 \geq \dots \geq \beta_n \geq 0\}.$$

The next lemma records the relation between random walks and linear subspaces intersecting Weyl chambers. The case $k = n$ of this lemma appeared in [6].

Lemma 2.3. *Let $x_1, \dots, x_n \in \mathbb{R}^d$ be arbitrary vectors and denote by $s_i = x_1 + \dots + x_i$, $1 \leq i \leq n$, their partial sums. Let $A : \mathbb{R}^n \rightarrow \mathbb{R}^d$ be the linear operator defined on the standard basis e_1, \dots, e_n of \mathbb{R}^n by $Ae_1 = x_1, \dots, Ae_n = x_n$. Then, the number of collections of indices $1 \leq i_1 < \dots < i_k \leq n$ such that $0 \in \text{Conv}(s_{i_1}, \dots, s_{i_k})$ is equal to the number of k -dimensional faces F in the convex cone C^B intersected non-trivially by the linear subspace $\text{Ker } A$.*

Proof. For a given collection of indices $1 \leq i_1 < \dots < i_k \leq n$ we have $0 \in \text{Conv}(s_{i_1}, \dots, s_{i_k})$ if and only if there exist $\alpha_1, \dots, \alpha_k \geq 0$ (not all of them being 0) such that $\alpha_1 s_{i_1} + \dots + \alpha_k s_{i_k} = 0$, or, equivalently,

$$\alpha_1(x_1 + \dots + x_{i_1}) + \alpha_2(x_1 + \dots + x_{i_2}) + \dots + \alpha_k(x_1 + \dots + x_{i_k}) = 0.$$

Rearranging the terms, we rewrite this condition as

$$\begin{aligned} x_1(\alpha_1 + \dots + \alpha_k) + \dots + x_{i_1}(\alpha_1 + \dots + \alpha_k) \\ + x_{i_1+1}(\alpha_2 + \dots + \alpha_k) + \dots + x_{i_2}(\alpha_2 + \dots + \alpha_k) \\ + \dots + x_{i_{k-1}+1}\alpha_k + \dots + x_{i_k}\alpha_k = 0. \end{aligned} \quad (25)$$

Introducing the new variables β_1, \dots, β_n as the coefficients of x_1, \dots, x_n , that is

$$\begin{aligned} \beta_1 = \dots = \beta_{i_1} &:= \alpha_1 + \dots + \alpha_k, \\ \beta_{i_1+1} = \dots = \beta_{i_2} &:= \alpha_2 + \dots + \alpha_k, \\ \dots, \\ \beta_{i_{k-1}+1} = \dots = \beta_{i_k} &:= \alpha_k, \\ \beta_{i_k+1} = \dots = \beta_n &:= 0, \end{aligned}$$

we can rewrite (25) as $\beta_1 x_1 + \dots + \beta_n x_n = 0$ or, equivalently, $(\beta_1, \dots, \beta_n) \in \text{Ker } A$. Our conditions on the α_i 's translate into the following equivalent condition on the β_i 's:

$$\beta_1 = \dots = \beta_{i_1} \geq \beta_{i_1+1} = \dots = \beta_{i_2} \geq \dots \geq \beta_{i_{k-1}+1} = \dots = \beta_{i_k} \geq \beta_{i_k+1} = \dots = \beta_n = 0, \quad (26)$$

where at least one inequality should be strict, that is $(\beta_1, \dots, \beta_n) \neq 0$. If $i_k = n$, then there are no β_i 's required to vanish. Summarizing, we have $0 \in \text{Conv}(s_{i_1}, \dots, s_{i_k})$ if and only if $F \cap \text{Ker } A \neq \{0\}$, where $F \subset \mathbb{R}^n$ is the closed convex cone defined by (26). Since any such F is a k -dimensional face of the convex cone C^B and, conversely, any k -face has this form, we obtain the required statement. \square

Proof of Theorem 1.2 given Theorem 2.1. Let $A : \mathbb{R}^n \rightarrow \mathbb{R}^d$ be the random linear operator defined on the standard basis e_1, \dots, e_n of \mathbb{R}^n by $Ae_1 = \xi_1, \dots, Ae_n = \xi_n$. By Lemma 2.3,

$$M_{n,k}^{(d)} = \frac{1}{2} \sum_{1 \leq i_1 < \dots < i_k \leq n} \mathbb{1}_{\{0 \notin \text{Conv}(S_{i_1}, \dots, S_{i_k})\}} = \frac{1}{2} \sum_{F \in \mathcal{F}_k(C^B)} \mathbb{1}_{\{F \cap \text{Ker } A = \{0\}\}} \quad \text{a.s.}, \quad (27)$$

where we used that both sums have the same number of terms $\#\mathcal{F}_k(C^B) = \binom{n}{k}$. Let us show that together with symmetric exchangeability assumption ($\pm\text{Ex}$), this implies that for every closed Weyl chamber $C_{\varepsilon, \sigma}^B$,

$$M_{n,k}^{(d)} \stackrel{d}{=} \frac{1}{2} \sum_{F \in \mathcal{F}_k(C_{\varepsilon, \sigma}^B)} \mathbb{1}_{\{F \cap \text{Ker } A = \{0\}\}}. \quad (28)$$

First note that $C_{\varepsilon, \sigma}^B = g(C^B)$ for $g := g_{\varepsilon((\bar{\sigma})^{-1}), (\bar{\sigma})^{-1}}$, where the permutation $\bar{\sigma}$ is defined by $\bar{\sigma} = (\sigma(n), \dots, \sigma(1))$. This holds by the fact that $ge_k = \varepsilon_{\bar{\sigma}(k)} e_{\bar{\sigma}(k)}$ (where $1 \leq k \leq n$), which ensures that $\bar{\sigma}$ arranges the absolute values of coordinates of points in $g(C^B)$ in an increasing order, so σ arranges them in a decreasing order as needed for $C_{\varepsilon, \sigma}^B$. Further, for any k -face $F \in \mathcal{F}_k(C^B)$,

$$\{g(F) \cap \text{Ker } A = \{0\}\} = \{F \cap g^{-1}(\text{Ker } A) = \{0\}\} = \{F \cap \text{Ker}(Ag) = \{0\}\}.$$

Since the random linear operator Ag satisfies $(Ag)e_1 = \varepsilon_{\sigma(n)} e_{\sigma(n)}, \dots, (Ag)e_n = \varepsilon_{\sigma(1)} e_{\sigma(1)}$, from ($\pm\text{Ex}$) it follows that $\text{Ker}(Ag) \stackrel{d}{=} \text{Ker } A$. Hence (28) follows from (27) as required.

Taking the expectation in (28) and then the mean over all $2^n n!$ pairs (ε, σ) , we obtain

$$\begin{aligned} 2\mathbb{E}M_{n,k}^{(d)} &= \mathbb{E} \left[\frac{1}{2^n n!} \sum_{\varepsilon \in \{-1, +1\}^n} \sum_{\sigma \in \text{Sym}(n)} \sum_{F \in \mathcal{F}_k(C_{\varepsilon, \sigma}^B)} \mathbb{1}_{\{F \cap \text{Ker } A = \{0\}\}} \right] \\ &= 2 \binom{n}{k} \frac{B(k, d-1) + B(k, d-3) + \dots}{2^k k!}, \end{aligned}$$

where the second equality is by Theorem 2.1 applied to $L := \text{Ker } A$. The fact that with probability one, $\text{Ker } A$ has codimension d and is in general position with respect to the reflection arrangement of type B_n is proved in [6, Lemma 6.3]. \square

We finish this section with the following observation discussed in the Introduction and related to Corollary 2.2. In the case $d = 1$, the random subspace $\text{Ker } A$ is a hyperplane a.s. It is easy to see that N_n , the number of positive terms of the random walk $(S_i)_{i=1}^n$, is a.s. equal to the number of vertices of the simplex

$$\{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : 1 \geq \beta_1 \geq \beta_2 \geq \dots \geq \beta_n \geq 0\}$$

lying in the open half-space $\{\beta \in \mathbb{R}^n : A \cdot \beta > 0\}$ with the boundary $\text{Ker } A$.

2.4. Proof of Theorem 2.1. In the special case $n = k$ Theorem 2.1 gives a formula for the number of Weyl chambers of type B_n intersected non-trivially by a linear subspace of codimension d in general position. It was established in [6] using the theory of hyperplane arrangements. Let us state this result.

Theorem 2.4 ([6]). *Let $L \subset \mathbb{R}^n$ be a deterministic linear subspace of codimension d in general position with respect to the reflection arrangement of type B_n . Let Q be sampled randomly and uniformly among the $2^n n!$ closed Weyl chambers $C_{\varepsilon, \sigma}^B$ of type B_n . Then,*

$$\begin{aligned} \mathbb{P}[L \cap Q = \{0\}] &\stackrel{def}{=} \frac{1}{2^n n!} \sum_{\varepsilon \in \{-1, +1\}^n} \sum_{\sigma \in \text{Sym}(n)} \mathbb{1}_{\{L \cap C_{\varepsilon, \sigma}^B = \{0\}\}} \\ &= \frac{2(B(n, d-1) + B(n, d-3) + \dots)}{2^n n!}. \end{aligned}$$

The following combinatorial proof deduces the assertion of Theorem 2.1 for general k from Theorem 2.4 without using any additional tools.

Enumeration of the k -faces of Weyl chambers of type B_n . Before proceeding to the proof, we introduce some notation and give few combinatorial examples. Recall that the k -dimensional faces of the Weyl chamber $C_{\varepsilon, \sigma}^B$ are denoted by $C_{\varepsilon, \sigma}^B(i_1, \dots, i_k)$; see (22). It is important to stress that the k -face $C_{\varepsilon, \sigma}^B(i_1, \dots, i_k)$ may be a k -face of another Weyl chamber $C_{\varepsilon', \sigma'}^B$ with some $(\varepsilon', \sigma') \neq (\varepsilon, \sigma)$.

Example 2.5. Consider the case $n = 8$, $k = 3$ and the convex cone given by the following set of conditions:

$$\underbrace{-\beta_2 = \beta_4}_{\text{group 1}} \geq \underbrace{\beta_1 = -\beta_6}_{\text{group 2}} \geq \underbrace{\beta_3}_{\text{group 3}} \geq \underbrace{\beta_5 = \beta_7 = \beta_8}_{\text{group 4}} = 0. \quad (29)$$

This cone is a 3-dimensional face of the Weyl chamber

$$-\beta_2 \geq \beta_4 \geq \beta_1 \geq -\beta_6 \geq \beta_3 \geq \beta_5 \geq \beta_7 \geq \beta_8 \geq 0. \quad (30)$$

However, it is also a 3-face of

$$\beta_4 \geq -\beta_2 \geq -\beta_6 \geq \beta_1 \geq \beta_3 \geq -\beta_8 \geq -\beta_5 \geq \beta_7 \geq 0$$

and, more generally, any of the chambers obtained from (30) by permuting the β 's inside the groups $(-\beta_2, \beta_4)$, $(\beta_1, -\beta_6)$, β_3 , $(\beta_5, \beta_7, \beta_8)$, and by changing any number of signs in the last group. The total number of such chambers is $2!2!1!3!2^3$.

We now introduce an enumeration of all k -faces of all Weyl chambers such that each face is counted exactly once. Let $\mathcal{P}_{n,k}$ be the set of all pairs (I, η) , where $I = (I_1, \dots, I_{k+1})$ is a partition of the set $\{1, \dots, n\}$ into $k+1$ disjoint distinguishable subsets ("groups") such that I_1, \dots, I_k are non-empty, whereas I_{k+1} may be empty or not, and $\eta : I_1 \cup \dots \cup I_k \rightarrow \{-1, +1\}$. We shall write $\eta_i := \eta(i)$. Given a pair $(I, \eta) \in \mathcal{P}_{n,k}$ define a closed k -dimensional convex cone

$$\begin{aligned} Q_{I, \eta} &:= \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \\ &\text{for all } 1 \leq l_1 \leq l_2 \leq k \text{ and } i_1 \in I_{l_1}, i_2 \in I_{l_2} \text{ we have } \eta_{i_1} \beta_{i_1} \geq \eta_{i_2} \beta_{i_2} \geq 0; \\ &\text{for all } i \in I_{k+1} \text{ we have } \beta_i = 0\}. \end{aligned}$$

As a consequence of these conditions, for all $1 \leq l \leq k$ and $i_1, i_2 \in I_l$ we have $\eta_{i_1} \beta_{i_1} = \eta_{i_2} \beta_{i_2}$.

Example 2.6. If $n = 8$, $k = 3$ and the partition I is given by $I_1 = \{2, 4\}$, $I_2 = \{1, 6\}$, $I_3 = \{3\}$, $I_4 = \{5, 7, 8\}$, and the signs are $\eta_1 = \eta_3 = \eta_4 = +1$, $\eta_2 = \eta_6 = -1$, then the cone $Q_{I,\eta}$ is given by the set of inequalities (29).

Given $(I, \eta) \in \mathcal{P}_{n,k}$ denote by $V_{I,\eta}$ the k -dimensional linear subspace of \mathbb{R}^n spanned by $Q_{I,\eta}$, that is

$$V_{I,\eta} := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \\ \text{for all } 1 \leq l \leq k \text{ and } i_1, i_2 \in I_l \text{ we have } \eta_{i_1} \beta_{i_1} = \eta_{i_2} \beta_{i_2}; \\ \text{for all } i \in I_{k+1} \text{ we have } \beta_i = 0\}.$$

Using $\gamma_l := \eta_i \beta_i$, where $i \in I_l$ is arbitrary and $l = 1, \dots, k$, as coordinates on $V_{I,\eta}$ allows us to identify this linear space with \mathbb{R}^k . There is a natural decomposition of $V_{I,\eta}$ into $2^k k!$ Weyl chambers of type B_k which have the form

$$V_{I,\eta}(\zeta, \tau) = \{(\beta_1, \dots, \beta_n) \in V_{I,\eta} : \zeta_1 \gamma_{\tau(1)} \geq \dots \geq \zeta_k \gamma_{\tau(k)} \geq 0\},$$

where $\zeta \in \{-1, +1\}^k$, $\tau \in \text{Sym}(k)$. One of these chambers, corresponding to $\zeta_i = +1$, $\tau(i) = i$ for all $1 \leq i \leq k$, is $Q_{I,\eta}$.

Example 2.7. If the pair (I, η) is the same as in Example 2.6, then the linear subspace $V_{I,\eta}$ is given by the following set of conditions

$$\underbrace{\gamma_1 := -\beta_2 = \beta_4}_{\text{group 1}} \in \mathbb{R}, \quad \underbrace{\gamma_2 := \beta_1 = -\beta_6}_{\text{group 2}} \in \mathbb{R}, \quad \underbrace{\gamma_3 := \beta_3}_{\text{group 3}} \in \mathbb{R}, \quad \underbrace{\beta_5 = \beta_7 = \beta_8}_{\text{group 4}} = 0. \quad (31)$$

It should be stressed that the linear subspaces $V_{I,\eta}$ are not pairwise different. For example, the set of conditions (31) is clearly equivalent to the following one:

$$\underbrace{-\beta_3}_{\text{former group 3}} \in \mathbb{R}, \quad \underbrace{\beta_1 = -\beta_6}_{\text{former group 2}} \in \mathbb{R}, \quad \underbrace{\beta_2 = -\beta_4}_{\text{former group 1}} \in \mathbb{R}, \quad \underbrace{\beta_5 = \beta_7 = \beta_8}_{\text{group 4}} = 0.$$

More generally, we can interchange the first k groups in an arbitrary way and multiply any number of groups by ± 1 , giving a total number of $2^k k!$ possibilities. Clearly, the cone $Q_{I,\eta}$ given by (29) is one of the $2^k k!$ Weyl chambers which constitute $V_{I,\eta}$. However, as was explained above, there are other pairs (I', η') such that $V_{I,\eta} = V_{I',\eta'}$ and $Q_{I,\eta}$ coincides with one of the chambers $V_{I',\eta'}(\zeta, \tau)$.

Proof of Theorem 2.1. The cones $Q_{I,\eta}$, where $(I, \eta) \in \mathcal{P}_{n,k}$, are pairwise different and exhaust all k -dimensional faces of the Weyl chambers of type B_n . The cone $Q_{I,\eta}$ belongs to $(\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}}$ Weyl chambers because, as was explained in Example 2.5, in the definition of the Weyl chamber containing $Q_{I,\eta}$ we can postulate any order of the elements $\eta_i \beta_i$, $i \in I_l$, for all $1 \leq l \leq k$, and, additionally, we can postulate any order of the elements $\pm \beta_i$, $i \in I_{k+1}$, with arbitrary chosen signs. It follows that

$$\sum_{\varepsilon \in \{-1, +1\}^n} \sum_{\sigma \in \text{Sym}(n)} \sum_{F \in \mathcal{F}_k(C_{\varepsilon, \sigma}^B)} \mathbb{1}_{\{F \cap L = \{0\}\}} = \sum_{(I, \eta) \in \mathcal{P}_{n,k}} (\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}} \mathbb{1}_{\{Q_{I,\eta} \cap L = \{0\}\}}. \quad (32)$$

In the rest of the proof we compute the right-hand side of (32). We may suppose that $k \geq d$ because otherwise Theorem 2.1 becomes trivial (L intersects all k -faces trivially).

Then, the codimension of $L \cap V_{I,\eta}$ in $V_{I,\eta}$ (which is an element of $\mathcal{L}(B_n)$) is d because L is in general position with respect to the reflection arrangement of type B_n in \mathbb{R}^n ; see (24). Also, $L \cap V_{I,\eta}$ is in general position with respect to the reflection arrangement of type B_k in $V_{I,\eta}$, as can be checked using the definition. It follows from Theorem 2.4 applied to the linear subspace $L \cap V_{I,\eta} \subset V_{I,\eta}$ that

$$\sum_{\zeta \in \{-1, +1\}^k} \sum_{\tau \in \text{Sym}(k)} \mathbb{1}_{\{L \cap V_{I,\eta}(\zeta, \tau) = \{0\}\}} = 2(B(k, d-1) + B(k, d-3) + \dots).$$

Multiplying this equality by $(\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}}$ and taking the sum over all $(I, \eta) \in \mathcal{P}_{n,k}$ we obtain

$$\begin{aligned} \sum_{(I, \eta) \in \mathcal{P}_{n,k}} \sum_{\zeta \in \{-1, +1\}^k} \sum_{\tau \in \text{Sym}(k)} (\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}} \mathbb{1}_{\{L \cap V_{I,\eta}(\zeta, \tau) = \{0\}\}} \\ = 2(B(k, d-1) + B(k, d-3) + \dots) \sum_{(I, \eta) \in \mathcal{P}_{n,k}} (\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}}. \end{aligned} \quad (33)$$

Let us look at the triple sum on the left-hand side of (33). Since any convex cone $Q_{I,\eta}$ can be represented in the form $V_{I',\eta'}(\zeta', \tau')$ in $2^k k!$ different ways, we have

$$\text{LHS}(33) = 2^k k! \sum_{(I, \eta) \in \mathcal{P}_{n,k}} (\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}} \mathbb{1}_{\{Q_{I,\eta} \cap L = \{0\}\}}. \quad (34)$$

Let us now compute the sum on the right-hand side of (33). For $j_1, \dots, j_k \in \mathbb{N}$ such that $j_1 + \dots + j_k \leq n$ denote by $\mathcal{P}_{n,k}(j_1, \dots, j_k)$ the set of all pairs $(I, \eta) \in \mathcal{P}_{n,k}$ such that $\#I_1 = j_1, \dots, \#I_k = j_k$. Let $j_{k+1} = n - j_1 - \dots - j_k$. The number of elements in $\mathcal{P}_{n,k}(j_1, \dots, j_k)$ is given by

$$\#\mathcal{P}_{n,k}(j_1, \dots, j_k) = \frac{n! 2^{j_1 + \dots + j_k}}{j_1! \dots j_{k+1}!}.$$

It follows that

$$\begin{aligned} \sum_{(I, \eta) \in \mathcal{P}_{n,k}} (\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}} &= \sum_{\substack{j_1, \dots, j_k \in \mathbb{N} \\ j_1 + \dots + j_k \leq n}} \sum_{(I, \eta) \in \mathcal{P}_{n,k}(j_1, \dots, j_k)} (\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}} \\ &= \sum_{\substack{j_1, \dots, j_k \in \mathbb{N} \\ j_1 + \dots + j_k \leq n}} \frac{n! 2^{j_1 + \dots + j_k}}{j_1! \dots j_{k+1}!} \cdot j_1! \dots j_{k+1}! 2^{n - (j_1 + \dots + j_k)} \\ &= \sum_{\substack{j_1, \dots, j_k \in \mathbb{N} \\ j_1 + \dots + j_k \leq n}} n! 2^n \\ &= n! 2^n \binom{n}{k}. \end{aligned} \quad (35)$$

Taking (33), (34), (35) together yields

$$\sum_{(I,\eta) \in \mathcal{P}_{n,k}} (\#I_1)! \dots (\#I_{k+1})! 2^{\#I_{k+1}} \mathbb{1}_{\{Q_{I,\eta} \cap L = \{0\}\}} = 2^n n! \binom{n}{k} \frac{2(B(k, d-1) + B(k, d-3) + \dots)}{2^k k!}.$$

In view of (32), this completes the proof of Theorem 2.1. \square

2.5. The reflection group and Weyl chambers of type A_{n-1} . We proceed to a geometric interpretation of Theorem 1.8. In the same way as random walks are related to the reflection group of type B_n , random bridges are related to the reflection group of type A_{n-1} . We start by recalling some definitions.

The *reflection group* $\mathcal{G}(A_{n-1})$ is the symmetric group $\text{Sym}(n)$ acting on \mathbb{R}^n by permuting the coordinates in an arbitrary way. The number of elements in $\mathcal{G}(A_{n-1})$ is $n!$. Note that this group leaves the hyperplane

$$L_0 := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \beta_1 + \dots + \beta_n = 0\}$$

invariant. The *fundamental Weyl chamber of type A_{n-1}* is the convex cone

$$\mathcal{C}(A_{n-1}) := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \beta_1 < \dots < \beta_n\}.$$

Acting on the closure of this cone by the elements of the group $\mathcal{G}(A_{n-1})$, we obtain the closed Weyl chambers of type A_{n-1} given by

$$C_\sigma^A := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \beta_{\sigma(1)} \geq \dots \geq \beta_{\sigma(n)}\}, \quad (36)$$

where $\sigma \in \text{Sym}(n)$. Note that $\cup_{\sigma \in \text{Sym}(n)} C_\sigma^A = \mathbb{R}^n$, and the interiors of these convex cones are disjoint. The *reflection arrangement $\mathcal{A}(A_{n-1})$ of type A_{n-1}* consists of the hyperplanes

$$\{\beta_i = \beta_j\}, \quad 1 \leq i < j \leq n. \quad (37)$$

The lattice $\mathcal{L}(A_{n-1})$ generated by the reflection arrangement $\mathcal{A}(A_{n-1})$ and the notion of general position are defined in the same way as in the B_n -case. The next result is an analogue of Theorem 2.1 for Weyl chambers of type A_{n-1} .

Theorem 2.8. *Let $L \subset \mathbb{R}^n$ be a deterministic linear subspace of codimension d in general position with respect to the reflection arrangement (37) of type A_{n-1} . Let Q be sampled randomly and uniformly among the $n!$ closed Weyl chambers C_σ^A of type A_{n-1} . Then, the expected number of k -dimensional faces of Q intersected by L in a trivial way is given by*

$$\begin{aligned} \mathbb{E} \left[\sum_{F \in \mathcal{F}_k(Q)} \mathbb{1}_{\{F \cap L = \{0\}\}} \right] &\stackrel{\text{def}}{=} \frac{1}{n!} \sum_{\sigma \in \text{Sym}(n)} \sum_{F \in \mathcal{F}_k(C_\sigma^A)} \mathbb{1}_{\{F \cap L = \{0\}\}} \\ &= \frac{2}{k!} \binom{n-1}{k-1} \left(\left[\begin{matrix} k \\ d-1 \end{matrix} \right] + \left[\begin{matrix} k \\ d-3 \end{matrix} \right] + \dots \right), \end{aligned}$$

where the $\left[\begin{matrix} k \\ j \end{matrix} \right]$'s are the Stirling numbers of the first kind defined in Theorem 1.8.

In the special case $k = n$, Theorem 2.8 reduces to the formula, which was proved in [6]:

$$\mathbb{P}[L \cap Q = \{0\}] \stackrel{\text{def}}{=} \frac{1}{n!} \sum_{\sigma \in \text{Sym}(n)} \mathbb{1}_{\{L \cap C_{\sigma}^A = \{0\}\}} = \frac{2}{n!} \left(\binom{n}{d-1} + \binom{n}{d-3} + \dots \right). \quad (38)$$

2.6. Proof of Theorem 1.8 given Theorem 2.8. We shall need a short notation for one of the closed Weyl chambers of type A_{n-1} :

$$C^A := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \beta_1 \geq \dots \geq \beta_n\}.$$

The next lemma is an analogue of Lemma 2.3. The case $k = n$ of this lemma appeared in [6].

Lemma 2.9. *Let $x_1, \dots, x_n \in \mathbb{R}^d$, where $n \geq 2$, be arbitrary vectors such that $x_1 + \dots + x_n = 0$. Denote by $s_i = x_1 + \dots + x_i$, $1 \leq i \leq n$, their partial sums. Let $A : \mathbb{R}^n \rightarrow \mathbb{R}^d$ be a linear operator defined on the standard basis e_1, \dots, e_n of \mathbb{R}^n by $Ae_1 = x_1, \dots, Ae_n = x_n$. Then the number of collections $1 \leq i_1 < \dots < i_k \leq n-1$ such that $0 \in \text{Conv}(s_{i_1}, \dots, s_{i_k})$ is equal to the number of $(k+1)$ -dimensional faces F of the convex cone C^A intersected non-trivially by the linear subspace $L_0 \cap \text{Ker } A$.*

Proof. For a given collection of indices $1 \leq i_1 < \dots < i_k \leq n-1$ we have $0 \in \text{Conv}(s_{i_1}, \dots, s_{i_k})$ if and only if there exist $\alpha_1, \dots, \alpha_k \geq 0$ (not all of them being 0) such that $\alpha_1 s_{i_1} + \dots + \alpha_k s_{i_k} = 0$, or, equivalently,

$$\alpha_1(x_1 + \dots + x_{i_1}) + \alpha_2(x_1 + \dots + x_{i_2}) + \dots + \alpha_k(x_1 + \dots + x_{i_k}) = 0.$$

After rearranging the terms, we can rewrite this condition as $\beta_1 x_1 + \dots + \beta_n x_n = 0$, where

$$\begin{aligned} \beta_1 &= \dots = \beta_{i_1} := \alpha_1 + \dots + \alpha_k - b, \\ \beta_{i_1+1} &= \dots = \beta_{i_2} := \alpha_2 + \dots + \alpha_k - b, \\ &\dots, \\ \beta_{i_{k-1}+1} &= \dots = \beta_{i_k} := \alpha_k - b, \\ \beta_{i_k+1} &= \dots = \beta_n := -b, \end{aligned}$$

and $b \in \mathbb{R}$ can be arbitrary due to the assumption $x_1 + \dots + x_n = 0$. Choose $b := \frac{1}{n}(i_1 \alpha_1 + \dots + i_k \alpha_k)$, which ensures that $\beta_1 + \dots + \beta_n = 0$. Our conditions on the α_i 's translate into the following equivalent conditions on the β_i 's:

$$\beta_1 = \dots = \beta_{i_1} \geq \beta_{i_1+1} = \dots = \beta_{i_2} \geq \dots \geq \beta_{i_{k-1}+1} = \dots = \beta_{i_k} \geq \beta_{i_k+1} = \dots = \beta_n, \quad (39)$$

$$\beta_1 + \dots + \beta_n = 0, \quad (40)$$

where at least one inequality in (39) should be strict, i.e. $(\beta_1, \dots, \beta_n) \neq 0$. That is, we have $0 \in \text{Conv}(s_{i_1}, \dots, s_{i_k})$ if and only if $F \cap L_0 \cap \text{Ker } A \neq \{0\}$, where $F \subset \mathbb{R}^n$ is the closed convex cone defined by (39). Since any such F is a $(k+1)$ -dimensional face of the Weyl chamber C^A and, conversely, any $(k+1)$ -face has this form, we obtain the required statement. \square

Proof of Theorem 1.8 given Theorem 2.8. Let $A : \mathbb{R}^n \rightarrow \mathbb{R}^d$ be a random linear operator defined on the standard basis e_1, \dots, e_n of \mathbb{R}^n by $Ae_1 = \xi_1, \dots, Ae_n = \xi_n$. By Lemma 2.9,

$$M_{n,k}^{(d)} = \frac{1}{2} \sum_{1 \leq i_1 < \dots < i_k \leq n} \mathbb{1}_{\{0 \notin \text{Conv}(s_{i_1}, \dots, s_{i_k})\}} = \frac{1}{2} \sum_{F \in \mathcal{F}_{k+1}(C^A)} \mathbb{1}_{\{F \cap L_0 \cap \text{Ker } A = \{0\}\}} \quad \text{a.s.},$$

where we also used that $\#\mathcal{F}_{k+1}(C^A) = \binom{n-1}{k}$. By the exchangeability assumption (Ex), for every closed Weyl chamber C_σ^A ,

$$M_{n,k}^{(d)} \stackrel{d}{=} \frac{1}{2} \sum_{F \in \mathcal{F}_{k+1}(C_\sigma^A)} \mathbb{1}_{\{F \cap L_0 \cap \text{Ker } A = \{0\}\}}.$$

Taking the expectation and then the mean over all $n!$ permutations $\sigma \in \text{Sym}(n)$, we obtain

$$\begin{aligned} 2\mathbb{E}M_{n,k}^{(d)} &= \mathbb{E} \left[\frac{1}{n!} \sum_{\sigma \in \text{Sym}(n)} \sum_{F \in \mathcal{F}_{k+1}(C_\sigma^A)} \mathbb{1}_{\{F \cap L_0 \cap \text{Ker } A = \{0\}\}} \right] \\ &= \frac{2}{(k+1)!} \binom{n-1}{k} \left(\binom{k+1}{d} + \binom{k+1}{d-2} + \dots \right), \end{aligned}$$

where the second equality is by Theorem 2.8 applied to the linear subspace $L := L_0 \cap \text{Ker } A$. The fact that with probability one, the linear subspace $L_0 \cap \text{Ker } A$ is in general position with respect to the hyperplane arrangement $\mathcal{A}(A_{n-1})$ and the codimension of this subspace in \mathbb{R}^n is $d+1$ is proved in [6, Lemma 6.2]. \square

2.7. Proof of Theorem 2.8. The proof is similar to that of Theorem 2.1, but several simplifications are possible.

Enumeration of the k -faces of Weyl chambers of type B_n . Again, we start with notation and examples. The k -dimensional faces of the Weyl chamber C_σ^A , see (36), are enumerated by collections $1 \leq i_1 < \dots < i_{k-1} \leq n-1$ as follows:

$$\begin{aligned} C_\sigma^A(i_1, \dots, i_{k-1}) &:= \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \beta_{\sigma(1)} = \dots = \beta_{\sigma(i_1)} \\ &\geq \beta_{\sigma(i_1+1)} = \dots = \beta_{\sigma(i_2)} \geq \dots \geq \beta_{\sigma(i_{k-1}+1)} = \dots = \beta_{\sigma(n)}\}. \end{aligned} \quad (41)$$

The next example shows that the k -face $C_\sigma^A(i_1, \dots, i_{k-1})$ may be a k -face of another Weyl chamber $C_{\sigma'}^A$ with some $\sigma' \neq \sigma$.

Example 2.10. Consider the case $n = 5$, $k = 3$ and the convex cone given by the following set of conditions:

$$\underbrace{\beta_2 = \beta_4}_{\text{group 1}} \geq \underbrace{\beta_1 = \beta_5}_{\text{group 2}} \geq \underbrace{\beta_3}_{\text{group 3}}. \quad (42)$$

This cone is a 3-dimensional face of the Weyl chamber

$$\beta_2 \geq \beta_4 \geq \beta_1 \geq \beta_5 \geq \beta_3. \quad (43)$$

However, it is also a 3-face of

$$\beta_4 \geq \beta_2 \geq \beta_5 \geq \beta_1 \geq \beta_3.$$

and, more generally, any of the chambers obtained from (43) by permuting the β_i 's inside the groups (β_2, β_4) , (β_1, β_5) , β_3 . The total number of such chambers is $2!2!1!$.

Next we shall introduce a notation for all k -faces of all Weyl chambers such that each face is counted exactly once. Let $\mathcal{R}_{n,k}$ be the set of all partitions $I = (I_1, \dots, I_k)$ of the set $\{1, \dots, n\}$ into k disjoint non-empty distinguishable subsets I_1, \dots, I_k . Given a partition $I \in \mathcal{R}_{n,k}$, define the closed k -dimensional convex cone

$$Q_I := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \text{for all } 1 \leq l_1 \leq l_2 \leq k \text{ and } i_1 \in I_{l_1}, i_2 \in I_{l_2} \text{ we have } \beta_{i_1} \geq \beta_{i_2}\}.$$

It follows from these conditions that for all $1 \leq l \leq k$ and $i_1, i_2 \in I_l$ we have $\beta_{i_1} = \beta_{i_2}$.

Example 2.11. If $n = 5$, $k = 3$ and the partition I is given by $I_1 = \{2, 4\}$, $I_2 = \{1, 5\}$, $I_3 = \{3\}$, then the cone Q_I is given by the set of inequalities (42).

For a partition $I \in \mathcal{R}_{n,k}$, denote by W_I the k -dimensional linear subspace of \mathbb{R}^n spanned by Q_i , that is

$$W_I := \{(\beta_1, \dots, \beta_n) \in \mathbb{R}^n : \text{for all } 1 \leq l \leq k \text{ and } i_1, i_2 \in I_l \text{ we have } \beta_{i_1} = \beta_{i_2}\}.$$

Using $\gamma_l := \beta_i$, where $i \in I_l$ is arbitrary and $l = 1, \dots, k$, as coordinates on W_I allows us to identify this linear space with \mathbb{R}^k . There is a natural decomposition of W_I into $k!$ Weyl chambers of type A_{k-1} of the form

$$W_I(\tau) = \{(\beta_1, \dots, \beta_n) \in W_I : \gamma_{\tau(1)} \geq \dots \geq \gamma_{\tau(k)}\},$$

where $\tau \in \text{Sym}(k)$. One of these chambers, corresponding to the identity permutation $\tau(i) = i$ for all $1 \leq i \leq k$, is Q_I .

Example 2.12. If the partition I is the same as in Example 2.11, then the linear space W_I is given by the following set of conditions

$$\gamma_1 := \underbrace{\beta_2 = \beta_4}_{\text{group 1}} \in \mathbb{R}, \quad \gamma_2 := \underbrace{\beta_1 = \beta_5}_{\text{group 2}} \in \mathbb{R}, \quad \gamma_3 := \underbrace{\beta_3}_{\text{group 3}} \in \mathbb{R}. \quad (44)$$

The spaces W_I are not pairwise distinct. For example, the set of conditions (44) is equivalent to the following one:

$$\underbrace{\beta_3}_{\text{former group 3}} \in \mathbb{R}, \quad \underbrace{\beta_1 = \beta_5}_{\text{former group 2}} \in \mathbb{R}, \quad \underbrace{\beta_2 = \beta_4}_{\text{former group 1}} \in \mathbb{R}.$$

More generally, we can interchange the k groups of equal β_i 's in an arbitrary way, so that in the list W_I , $I \in \mathcal{R}_{n,k}$, each linear subspace appears $k!$ times under different names.

Proof of Theorem 2.8. The cones Q_I , where $I \in \mathcal{R}_{n,k}$, are pairwise distinct and exhaust all k -dimensional faces of the Weyl chambers of type A_{n-1} . The cone Q_I belongs to $(\#I_1)! \dots (\#I_k)!$ different Weyl chambers because in the definition of the Weyl chamber containing Q_I we can impose an arbitrary ordering of the elements β_i , $i \in I_l$, for all $1 \leq l \leq k$; see Example 2.10. It follows that

$$\sum_{\sigma \in \text{Sym}(n)} \sum_{F \in \mathcal{F}_k(C_\sigma^A)} \mathbb{1}_{\{F \cap L = \{0\}\}} = \sum_{I \in \mathcal{R}_{n,k}} (\#I_1)! \dots (\#I_k)! \mathbb{1}_{\{Q_I \cap L = \{0\}\}}. \quad (45)$$

In the rest of the proof we compute the right-hand side of (45). We may assume that $k \geq d + 1$ since otherwise Theorem 2.8 is trivial. Recall that the linear space L has codimension d in \mathbb{R}^n and is in general position with respect to the reflection arrangement

of type A_{n-1} in \mathbb{R}^n . It follows from the definition of the general position, see (24), that the linear subspace $L \cap W_I \subset W_I$ has codimension d in W_I and is in general position with respect to the reflection arrangement of type A_{k-1} in W_I . It follows from (38) applied to $L \cap W_I \subset W_I$ that

$$\sum_{\tau \in \text{Sym}(k)} \mathbb{1}_{\{L \cap W_I(\tau) = \{0\}\}} = 2 \left(\binom{k}{d-1} + \binom{k}{d-3} + \dots \right).$$

Multiplying this equality by $(\#I_1)! \dots (\#I_k)!$ and taking the sum over all partitions $I \in \mathcal{R}_{n,k}$, we obtain

$$\begin{aligned} \sum_{I \in \mathcal{R}_{n,k}} \sum_{\tau \in \text{Sym}(k)} (\#I_1)! \dots (\#I_k)! \mathbb{1}_{\{L \cap W_I(\tau) = \{0\}\}} \\ = 2 \left(\binom{k}{d-1} + \binom{k}{d-3} + \dots \right) \sum_{I \in \mathcal{R}_{n,k}} (\#I_1)! \dots (\#I_k)!. \end{aligned} \quad (46)$$

Since any k -face Q_I can be represented as $W_{I'}(\tau')$ in $k!$ ways, see Example 2.12, and the sets I_1, \dots, I_k are (up to their order) the same in all representations,

$$\text{LHS}(46) = k! \sum_{I \in \mathcal{R}_{n,k}} (\#I_1)! \dots (\#I_k)! \mathbb{1}_{\{Q_I \cap L = \{0\}\}}. \quad (47)$$

Let us now compute the sum on the right-hand side of (46). For $j_1, \dots, j_k \in \mathbb{N}$ such that $j_1 + \dots + j_k = n$ denote by $\mathcal{R}_{n,k}(j_1, \dots, j_k)$ the set of all partitions $I \in \mathcal{R}_{n,k}$ such that $\#I_1 = j_1, \dots, \#I_k = j_k$. The number of elements in $\mathcal{R}_{n,k}(j_1, \dots, j_k)$ is given by

$$\#\mathcal{R}_{n,k}(j_1, \dots, j_k) = \frac{n!}{j_1! \dots j_k!}.$$

It follows that

$$\begin{aligned} \sum_{I \in \mathcal{R}_{n,k}} (\#I_1)! \dots (\#I_k)! &= \sum_{\substack{j_1, \dots, j_k \in \mathbb{N} \\ j_1 + \dots + j_k = n}} \sum_{I \in \mathcal{R}_{n,k}(j_1, \dots, j_k)} (\#I_1)! \dots (\#I_k)! \\ &= \sum_{\substack{j_1, \dots, j_k \in \mathbb{N} \\ j_1 + \dots + j_k = n}} \frac{n!}{j_1! \dots j_k!} \cdot j_1! \dots j_k! \\ &= n! \binom{n-1}{k-1}. \end{aligned} \quad (48)$$

Taking (46), (47), (48) together yields

$$\sum_{I \in \mathcal{R}_{n,k}} (\#I_1)! \dots (\#I_k)! \mathbb{1}_{\{Q_I \cap L = \{0\}\}} = n! \binom{n-1}{k-1} \frac{2}{k!} \left(\binom{k}{d-1} + \binom{k}{d-3} + \dots \right).$$

In view of (45), this completes the proof of Theorem 2.8. \square

3. PROOFS BY REDUCTION TO NON-ABSORPTION PROBABILITY

In this section we give alternative proofs of Theorems 1.2 and 1.8.

3.1. Proof of Theorem 1.2. This proof rests on the following result obtained in [6], which is the special case of Theorem 1.2 with $k = n$:

Theorem 3.1 ([6]). *Let $(S_i)_{i=1}^n$ be a random walk in \mathbb{R}^d satisfying assumptions $(\pm Ex)$ and (GP) . Then*

$$\mathbb{P}[0 \notin \text{Conv}(S_1, \dots, S_n)] = \frac{2(B(n, d-1) + B(n, d-3) + \dots)}{2^n n!}.$$

We want to use this result to compute $\mathbb{P}[0 \notin \text{Conv}(S_{i_1}, \dots, S_{i_k})]$ but we cannot apply it directly because the increments $\xi'_1 := S_{i_1}, \dots, \xi'_k := S_{i_k} - S_{i_{k-1}}$ in general are not exchangeable. We restore the exchangeability by introducing an additional random reshuffling of the ξ'_i 's, which is the main idea of the following proof.

Proof of Theorem 1.2. Take some $1 \leq i_1 < \dots < i_k \leq n$ and let $i_0 := 0$. We subdivide the collection (ξ_1, \dots, ξ_n) into k groups of lengths

$$j_1 := i_1 > 0, \quad j_2 := i_2 - i_1 > 0, \quad \dots, \quad j_k := i_k - i_{k-1} > 0$$

and one group of length $n - (j_1 + \dots + j_k) = n - i_k \geq 0$ as follows:

$$\underbrace{(\xi_1, \dots, \xi_{i_1})}_{\text{length } j_1}, \quad \underbrace{(\xi_{i_1+1}, \dots, \xi_{i_2})}_{\text{length } j_2}, \quad \dots, \quad \underbrace{(\xi_{i_{k-1}+1}, \dots, \xi_{i_k})}_{\text{length } j_k}, \quad (\xi_{i_k+1}, \dots, \xi_n). \quad (49)$$

The last, $(k+1)$ -st group, will be mostly ignored in the sequel since the outcome of the corresponding terms in the definition of $M_{n,k}^{(d)}$ does not depend on these values.

Denote by ξ'_l the sum of the ξ_i 's in the l -th group above, that is

$$\xi'_l := \xi_{i_{l-1}+1} + \dots + \xi_{i_l}, \quad l = 1, \dots, k.$$

Let τ be a random permutation that is uniformly distributed on $\text{Sym}(k)$ and independent of (ξ_1, \dots, ξ_n) . We reshuffle the ξ'_i 's as follows:

$$\eta_i := \xi'_{\tau(i)}, \quad i = 1, \dots, k.$$

We claim that these k random vectors satisfy the symmetric exchangeability assumption $(\pm Ex)$ (with k substituted for n). The exchangeability is by the construction and the symmetry follows by the symmetry of the joint distribution of (ξ_1, \dots, ξ_n) . Let us give the details.

Let σ_1 be any permutation of the set $\{1, \dots, k\}$ and let $\varepsilon = (\varepsilon_1, \dots, \varepsilon_k) \in \{-1, +1\}^k$ be a sequence of ± 1 's of length k . By the fact that $\tau\sigma_1 \stackrel{d}{=} \tau$, we have the distributional identity

$$(\varepsilon_1 \eta_{\sigma_1(1)}, \dots, \varepsilon_k \eta_{\sigma_1(k)}) = (\varepsilon_1 \xi'_{\tau(\sigma_1(1))}, \dots, \varepsilon_k \xi'_{\tau(\sigma_1(k))}) \stackrel{d}{=} (\varepsilon_1 \eta_1, \dots, \varepsilon_k \eta_k),$$

from which we deduce

$$\mathbb{P}[(\varepsilon_1 \eta_{\sigma_1(1)}, \dots, \varepsilon_k \eta_{\sigma_1(k)}) \in \cdot] = \frac{1}{k!} \sum_{\sigma \in \text{Sym}(k)} \mathbb{P}[(\varepsilon_1 \xi'_{\sigma(1)}, \dots, \varepsilon_k \xi'_{\sigma(k)}) \in \cdot]. \quad (50)$$

Let $\sigma \in \text{Sym}(k)$ be any permutation of length k . Permute the first k groups of the list (49) according to σ and then multiply the reordered groups by $\varepsilon_1, \dots, \varepsilon_k$, respectively. The $(k+1)$ -st group stays unchanged. Clearly, our symmetric exchangeability assumption

($\pm\text{Ex}$) on (ξ_1, \dots, ξ_n) implies that this operation does not change the distribution. That is, ignoring the last group, we have the distributional equality

$$(\xi_1, \dots, \xi_{i_k}) \stackrel{d}{=} \underbrace{(\varepsilon_1 \xi_{i_{\sigma(1)}-1+1}, \dots, \varepsilon_1 \xi_{i_{\sigma(1)}})}_{\text{length } j_{\sigma(1)}} \underbrace{(\varepsilon_2 \xi_{i_{\sigma(2)}-1+1}, \dots, \varepsilon_2 \xi_{i_{\sigma(2)}})}_{\text{length } j_{\sigma(2)}} \dots \underbrace{(\varepsilon_k \xi_{i_{\sigma(k)}-1+1}, \dots, \varepsilon_k \xi_{i_{\sigma(k)}})}_{\text{length } j_{\sigma(k)}}.$$

Since the sum of random vectors in the l -th group is $\varepsilon_l \xi'_{\sigma(l)}$, we see that the distributions under the sum in (50) do not depend on $\varepsilon_1, \dots, \varepsilon_k$. This proves the stated symmetric exchangeability of (η_1, \dots, η_k) . This also gives the distributional identity

$$(S_{j_{\sigma(1)}}, S_{j_{\sigma(1)}+j_{\sigma(2)}}, \dots, S_{j_{\sigma(1)}+\dots+j_{\sigma(k)}}) \stackrel{d}{=} (\xi'_{\sigma(1)}, \xi'_{\sigma(1)} + \xi'_{\sigma(2)}, \dots, \xi'_{\sigma(1)} + \dots + \xi'_{\sigma(k)}). \quad (51)$$

Now introduce the new random walk

$$T_i := \eta_1 + \dots + \eta_i, \quad i = 1, \dots, k.$$

From (51) it follows that the random walk $(T_i)_{i=1}^k$ satisfies the general position assumption (GP) (with k substituted for n) since $(S_i)_{i=1}^n$ satisfies (GP). Then Theorem 3.1 applies to $(T_i)_{i=1}^k$, and by taking the mean over all $\sigma \in \text{Sym}(k)$ in (51), we obtain

$$\begin{aligned} \frac{1}{k!} \sum_{\sigma \in \text{Sym}(k)} \mathbb{P}[0 \notin \text{Conv}(S_{j_{\sigma(1)}}, S_{j_{\sigma(1)}+j_{\sigma(2)}}, \dots, S_{j_{\sigma(1)}+\dots+j_{\sigma(k)}})] \\ = \mathbb{P}[0 \notin \text{Conv}(T_1, \dots, T_k)] \\ = \frac{2(B(k, d-1) + B(k, d-3) + \dots)}{2^k k!}. \end{aligned} \quad (52)$$

Finally, sum the equations above over all tuples (j_1, \dots, j_k) from the following set:

$$J_{n,k} := \{(j_1, \dots, j_k) \in \mathbb{N}^k : j_1 + \dots + j_k \leq n\}.$$

Since the cardinality of $J_{n,k}$ is $\binom{n}{k}$ and the last expression in (52) does not depend on (j_1, \dots, j_k) , we obtain

$$\begin{aligned} \frac{2}{2^k k!} \binom{n}{k} (B(k, d-1) + B(k, d-3) + \dots) \\ = \frac{1}{k!} \sum_{\sigma \in \text{Sym}(k)} \sum_{(j_1, \dots, j_k) \in J_{n,k}} \mathbb{P}[0 \notin \text{Conv}(S_{j_{\sigma(1)}}, S_{j_{\sigma(1)}+j_{\sigma(2)}}, \dots, S_{j_{\sigma(1)}+\dots+j_{\sigma(k)}})] \\ = \sum_{(j_1, \dots, j_k) \in J_{n,k}} \mathbb{P}[0 \notin \text{Conv}(S_{j_1}, S_{j_1+j_2}, \dots, S_{j_1+\dots+j_k})] \\ = \sum_{1 \leq i_1 < \dots < i_k \leq n} \mathbb{P}[0 \notin \text{Conv}(S_{i_1}, S_{i_2}, \dots, S_{i_k})]. \end{aligned}$$

To complete the proof, observe that the right-hand side is nothing but $2\mathbb{E}M_{n,k}^{(d)}$; see (6). \square

3.2. Proof of Theorem 1.8. The proof is based on the following result obtained in [6]:

Theorem 3.2 ([6]). *Let $(S_i)_{i=1}^n$, $n \geq 2$, be a random bridge in \mathbb{R}^d satisfying assumptions (Br), (Ex), (GP'). Then,*

$$\mathbb{P}[0 \notin \text{Conv}(S_1, \dots, S_{n-1})] = \frac{2}{n!} \left(\binom{n}{d} + \binom{n}{d-2} + \dots \right).$$

Proof of Theorem 1.8. Take some $1 \leq i_1 < \dots < i_k \leq n-1$ and put $i_0 := 0$, $i_{k+1} := n$. We subdivide the collection (ξ_1, \dots, ξ_n) into $k+1$ groups of lengths

$$j_1 := i_1 > 0, \quad j_2 := i_2 - i_1 > 0, \quad \dots, \quad j_{k+1} := n - i_k > 0$$

as follows:

$$\underbrace{(\xi_1, \dots, \xi_{i_1})}_{\text{length } j_1}, \quad \underbrace{(\xi_{i_1+1}, \dots, \xi_{i_2})}_{\text{length } j_2}, \quad \dots, \quad \underbrace{(\xi_{i_k+1}, \dots, \xi_n)}_{\text{length } j_{k+1}}. \quad (53)$$

Denote by ξ'_l the sum of the ξ_i 's in the l -th group above, that is

$$\xi'_l := \xi_{i_{l-1}+1} + \dots + \xi_{i_l}, \quad l = 1, \dots, k+1.$$

Let τ be a random permutation that is uniformly distributed on $\text{Sym}(k+1)$ and independent with (ξ_1, \dots, ξ_n) . We reshuffle the ξ'_i 's as follows:

$$\eta_i := \xi'_{\tau(i)}, \quad i = 1, \dots, k+1.$$

By the construction, the $k+1$ random vectors η_i satisfy the exchangeability assumption (Ex) (with $k+1$ substituted for n).

Take any permutation σ of the set $\{1, \dots, k+1\}$ and permute the $k+1$ groups of the above list (53) according to σ . The exchangeability assumption (Ex) implies that this does not change the distribution of (ξ_1, \dots, ξ_n) , that is

$$(\xi_1, \dots, \xi_n) \stackrel{d}{=} \underbrace{(\xi_{i_{\sigma(1)}-1+1}, \dots, \xi_{i_{\sigma(1)}})}_{\text{length } j_{\sigma(1)}}, \quad \underbrace{(\xi_{i_{\sigma(2)}-1+1}, \dots, \xi_{i_{\sigma(2)}})}_{\text{length } j_{\sigma(2)}}, \quad \dots, \quad \underbrace{(\xi_{i_{\sigma(k+1)}-1+1}, \dots, \xi_{i_{\sigma(k+1)}})}_{\text{length } j_{\sigma(k+1)}}.$$

The sum of random vectors in the l -th group is $\xi'_{\sigma(l)}$, which gives the distributional identity

$$(S_{j_{\sigma(1)}}, S_{j_{\sigma(1)}+j_{\sigma(2)}}, \dots, S_{j_{\sigma(1)}+\dots+j_{\sigma(k+1)}}) \stackrel{d}{=} (\xi'_{\sigma(1)}, \xi'_{\sigma(1)} + \xi'_{\sigma(2)}, \dots, \xi'_{\sigma(1)} + \dots + \xi'_{\sigma(k+1)}), \quad (54)$$

which is analogous to (51). Note that the last vectors in both sides equal zero a.s. since $j_1 + \dots + j_{k+1} = n$ and $S_n = 0$ a.s.

Now introduce the partial sums

$$T_i := \eta_1 + \dots + \eta_i, \quad i = 1, \dots, k+1.$$

It follows from (54) that $T_{k+1} = S_n = 0$ a.s. and the random bridge $(T_i)_{i=1}^{k+1}$ satisfies the general position assumption (GP') (with $k+1$ substituted for n) since the random bridge

$(S_i)_{i=1}^n$ satisfies (GP'). Applying Theorem 3.2 to $(T_i)_{i=1}^{k+1}$, we obtain

$$\begin{aligned} \frac{1}{(k+1)!} \sum_{\sigma \in \text{Sym}(k+1)} \mathbb{P}[0 \notin \text{Conv}(S_{j_{\sigma(1)}}, S_{j_{\sigma(1)+j_{\sigma(2)}}}, \dots, S_{j_{\sigma(1)+\dots+j_{\sigma(k)}}})] \\ = \mathbb{P}[0 \notin \text{Conv}(T_1, \dots, T_k)] \\ = \frac{2}{(k+1)!} \left(\binom{k+1}{d} + \binom{k+1}{d-2} + \dots \right). \end{aligned} \quad (55)$$

Now we take the sum over all (j_1, \dots, j_{k+1}) from the set

$$J_{n,k}^* := \{(j_1, \dots, j_{k+1}) \in \mathbb{N}^{k+1} : j_1 + \dots + j_{k+1} = n\}.$$

Since the cardinality of $J_{n,k}^*$ is $\binom{n-1}{k}$ and the last expression in (55) does not depend on (j_1, \dots, j_{k+1}) , we obtain

$$\begin{aligned} \binom{n-1}{k} \frac{2}{(k+1)!} \left(\binom{k+1}{d} + \binom{k+1}{d-2} + \dots \right) \\ = \frac{1}{(k+1)!} \sum_{\sigma \in \text{Sym}(k+1)} \sum_{(j_1, \dots, j_{k+1}) \in J_{n,k}^*} \mathbb{P}[0 \notin \text{Conv}(S_{j_{\sigma(1)}}, S_{j_{\sigma(1)+j_{\sigma(2)}}}, \dots, S_{j_{\sigma(1)+\dots+j_{\sigma(k)}}})] \\ = \sum_{(j_1, \dots, j_{k+1}) \in J_{n,k}^*} \mathbb{P}[0 \notin \text{Conv}(S_{j_1}, S_{j_1+j_2}, \dots, S_{j_1+\dots+j_k})] \\ = \sum_{1 \leq i_1 < \dots < i_k \leq n-1} \mathbb{P}[0 \notin \text{Conv}(S_{i_1}, S_{i_2}, \dots, S_{i_k})]. \end{aligned}$$

To complete the proof, observe that the right-hand side is nothing but $2\mathbb{E}M_{n,k}^{(d)}$; see (19). \square

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